### The Evolution of Snp Petrom Stock List - Study Through Autoregressive Models

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#### **ABSTRACT**

Stock exchange market is one of the most dynamic and unpredictable markets. In this context, this work intends to analyze the SNP Petrom shares on the REGS market, based on the chronological series.

The economic series are often not stationary, but they can be stationarized by different data transformations. The simplest method used for stationarizing a series is to apply differentiating operators of various classes on the series. After applying this operator, a stationary series that can be modified by an ARIMA (p.q) process is usually obtained.

Most time series with economic content include a seasonal component besides the trend and random component.

The purpose of this work is to estimate the parameters of an ARIMA (p,d,q) model for SNP Petrom shares, where p is the number of autoregressive terms, d is the integration level of the series (how many times the series must be differentiated in order to become stationary) and q is the number of moving average terms (MA). **Key words:** list, economic series, autoregressive models

#### 1. INTRODUCTION

In literature the determination of the best ARIMA(p,d,q) sample in order to shape certain remarks for a series of time entails an assembly of techniques and methods, better known as the Box-Jenkins methodology.

A process  $\{Y_t\}$ , t belongs to Z, it admits a representation ARIMA(p,d,q) should this meet the subsequent equality:  $\Phi(L)(1-L)^dY_t=\Theta(L)\epsilon_t$ , whereas  $\epsilon_t$  is a white noise, the two polinomes  $\Phi(L)=1-\sum \phi_i L^i$ ,  $\Theta(L)=1-\sum \theta_i L^i$  have roots larger than one, as the initial conditions  $y_{-p-d}$ ,...

 $y_{-1}$ ,  $\varepsilon_{-q}$ , ...,  $\varepsilon_{-1}$  are not correlated with the random variables  $\varepsilon_0$ ,  $\varepsilon_1$ ,...,  $\varepsilon_t$ ,...

### 2. BUILDING THE MODEL WITH BOX-JENKINS METHODOLOGY

The Box-Jenkins methodology comprises three main aspects:

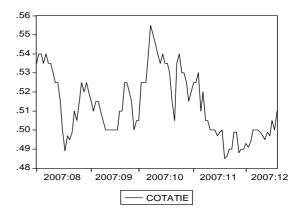
- identification;
- estimate:
- checking.

### Sample identification

Having available the sample of remarks on the evolution of SNP Petrom share quotation, a series of transformations must be brought to these so as to induce stationarity.

In case of time series describing the processes on the financial market, a scale transformation appears necessary, whereas most of the time the initial i series is being applied a logarithmic filter, in order to have a stationary series.

The next step is the elimination of the determinist component, after finding the possible oscillations present in the evolution of the series (Figure 1.).



**Figure 1** – Average price evolution of Petrom SA shares on the market

Currently we are able to determine for which values of the parameters p and q the ARMA(p,q) process shape to the best in the stationary series obtained. A criterion in this regard is the behaviour of the autocorrelation (ACF) and of	. ***** 1 0.67 0.02 4442 0.00   . .   3 6 5 .7 0 	
the partial autocorrelation (PACF) functions.	. .   4 3 8 .3 0	,
	. ***** 1 0.63 0.05 4861 0.00   . .   5 3 0 .6 0	)
Corelograma p_RRC	. ***** 1 0.61 0.04 5052 0.00	)
Included observations: 489	6 3 6 .1 0	
Partial	. ***** 1 0.59 0.00 5230 0.00   . .   7 2 5 .5 0	)
Autocorrel Correlati Q- ation on AC PAC Stat Prob	. **** 1 0.57 0.00 5397 0.00	)
. **** . *** 0.97 0.97 470. 0.00	. .   8 2 9 .3 0 . **** 1 0.55 0.00 5553 0.00	)
***	9 3 4 .5 0	
. ***** 0.95 0.12 916. 0.00	. **** 2 0.53 0.09 5702 0.00   . *   0 9 2 .0 0	)
	. **** 2 0.52 0.01 5843 0.00	)
. ***** 0.92 0.00 1338 0.00 **  . .   3 4 1 .0 0	. .   1 5 8 .2 0 . **** 2 0.51 0.07 5979 0.00	)
. ***** 0.90 0.04 1738 0.00	. *  2 4 9 .2 0	•
**  . .   4 0 8 .6 0	. **** 2 0.50 0.01 6110 0.00	)
. ***** 0.87 0.05 2117 0.00	. .   3 4 6 .1 0 . **** 2 0.49 0.02 6236 0.00	)
**  . .   5 4 4 .6 0	4 5 0 .8 0	
. *****	. **** 2 0.48 0.03 6358 0.00	)
	. .   5 5 0 .5 0 	)
. ****	. .   6 6 4 .8 0	
. ***** 0.79 0.07 3131 0.00	. **** 2 0.46 0.01 6588 0.00   . .   7 6 2 .7 0	)
*  * .  8 8 6 .5 0	· -	
. ***** 0.77 0.03 3429 0.00	. *** 2 0.45 0.05 6696 0.00   * .   8 5 8 .6 0	)
*  . .   9 1 1 .1 0	. *** 2 0.44 0.00 6799 0.00	)
. **** 1 0.74 0.01 3707 0.00 *  . .   0 5 1 .1 0	. ·***	)
. ***** 1 0.72 0.01 3967 0.00	. .   0 3 0 .1 0	
*  . .   1 0 8 .8 0	. *** 3 0.42 0.02 6991 0.00   . .   1 4 4 .2 0	
. ***** 1 0.69 0.00 4212 0.00	. *** 3 0.41 0.05 7082 0.00   . .   2 7 1 .5 0	)
2 7 5 .5 0	. *** . .   3 0.40 - 7170 0.00	)

		3	9	0.02	.7	0
الماد عاد عاد		2	0.40	9	7056	0.00
. ***		3	0.40	0.04	7256	0.00
	. .	4	3	5	.4	0
				-		
. ***		3	0.39	0.00	7339	0.00
	. .	5	7	5	.9	0
. ***		3	0.39	0.07	7422	0.00
	. *	6	4	5	.3	0

We can see that ACF decreases very slowly (up to 36 lags are statistically significant), as PACF dramatically decreases after the first lag. ACF suggests that the series of prices is not stationary, and it must be differentiated before applying the Box-Jenkins methodology. The test for the unit-root Dickey Fuller set out below proves that our series is actually integrated of order 1 (and not more).

## Null Hypothesis: P\_RRC has a unit root

**Exogenous: Constant** 

Lag Length: 1 (Automatic based on SIC, MAXLAG=17)

		t- Statistic Prob.*
Augme test statisti	nted Dickey-F	- Fuller 2.71968 5 0.0714
Test crit	tical 1% level	3.44355 1
	5% level	2.86725 5
	10% level	- 2.56987 6

<sup>\*</sup>MacKinnon (1996) one-sided p-values.

# **Augmented Dickey-Fuller Test Equation**

Dependent Variable: D(P\_RRC)

Method: Least Squares

Sample (adjusted): 1/04/2006

11/15/2007

Included observations: 487 after

adjustments

Variabl Coeffic Std. t- Prob.

P_RRC(-1	)6	059	5	8
D(P_RRC	(0.13041	0.044	2.89979	0.003
-1))	6	974	1	9
	0.00233	0.000	2.64845	0.008
С	3	881	0	4
	0.02903	Mean		-4.52E-
R-squared	8	depend	lent var	05
Adjusted R	-0.02502	S.D. de	ependent	0.0027
squared	6	var		04
				-
S.E. o	f 0.00267	Akaike	e info	9.00714
regression	0	criterio	n	1
				-
Sum squared	10.00345	Schwa	rz	8.98134
resid	1	criterio	n	1
Log	2196.23			7.2372
likelihood	9	F-statis	stic	91
Durbin-	2.00035	Prob(F	'-	0.0008
Watson stat	4	statistic	c)	00

Error Statistic

0.02463 0.009 2.71968 0.006

ient

e

## Null Hypothesis: D(P\_RRC) has a unit root

**Exogenous: Constant** 

Lag Length: 0 (Automatic based on SIC,

MAXLAG=17)

		t- Statistic	Prob.*
Augmented statistic	Dickey-Fuller	test 19.5410 9	0.0000
Test critica values:	l 1% level	3.44355 1	
	5% level	2.86725 5	
	10% level	2.56987 6	

<sup>\*</sup>MacKinnon (1996) one-sided p-values.

# **Augmented Dickey-Fuller Test Equation**

Dependent Variable: D(P\_RRC,2)

Method: Least Squares

Sample (adjusted): 1/04/2006

11/15/2007			. .	. .	7		0.06	11.5 23	0.11
Variable	Coeffic ient Std. Error	t- Prob Statistic .	. .	. .	8	5 0.03 9	_	12.2 83	•
D(P_RRC(- 1))	- 0.88085 8 0.045077	19.5410 0.00   9 00	. .	. .	9	0.00	0.01	12.2 83	0.19 8
C	-4.00E- 05 0.000122	- 0.32867 0.74   9 25	. .	. .	1	0.05 4	0.04 3	13.7 44	0.18 5
	Mean 0.44050 dependent		. .	. .	1 1	0.02 9	0.01 3	14.1 59	0.22 4
R-squared Adjusted	6 var S.D. R-0.43935 dependent	-2.05E-06	. .	. .	1 2	7	2	15.2 92	6
squared S.E.	2 var of 0.00268 Akaike info		. .	. .	1 3	0.01	0.02 0 -	15.3 38	0.28 7
resid Log	8 criterion red 0.00350 Schwarz 4 criterion 2192.54	-8.996081 -8.978881	* .   . .	* .	1 4 1 5	0.07 7 0.00 9	0.08 1 0.02 1	18.3 11 18.3 52	3
likelihood Durbin- Watson stat	6 F-statistic 1.99691 Prob(F- 4 statistic)	0.000000	. .	. .	1 6	- 0.00 9	- 0.01 3	92	0.30
integrated of	ng established that the	sted in ACF	. .	. .	1 7	0.00 7 -	0.00 4 -	18.4 19	0.36
Sample: 1/18/2008	for the first difference d(p 1/02/2006	5_RRC).	. .	. .	1 8	0.03 5	0.03 3 -	19.0 53	0.38 9
Included 488	observations:		* .	* .	1 9	0.10 7	0.09	24.8 85	0.16 4
Autocorrel ation	Correlation AC PAGE		. .	. .	2	0.02 4	0.01	25.1 82	0.19 5
. *	. * 0.11 0.11	1 6.96 0.00 97 8	* .	* .	2	0.09	0.08	29.4 18	0.10 4
. .	. . 0.00 0.01   2 2 2 -	1 6.97 0.03	. .	. .	2 2	0.01	0.00 5	29.4 80	0.13 2
* .	3 2 2	6 8.89 0.03 61 1 3 9.10 0.05	. .	. .		0.01	0.00 8	29.5 68	0.16 2
. .	4 1 6 .l. 0.02 0.01	67 8	. .	. .	2 4	0.01	0.00 6	29.6 38	0.19 7
. .	* . 0.04 0.05   6 9 9	5 10.5 0.10	. .   . .	. . . .	5	0.00 4 0.01	2	48	8

	. .	 	. .	6 2 7		6 0.02 7	16 30.0 96	0 0.31 0
	. .		. .	2 8	0.00 9	0.00	30.1 43	0.35 6
	. .		. .	2	0.01	0.01		0.40 4
	. .		. .	3	0.04	0.04 9		0.41 1
1	. .	1	* .	3	0.05 6	0.06 6	32.7 02	
I	. .	1	. .	3 2	0.03	0.04	33.2 60	0.40 6
I	. .	1	. .	3	0.00	0.03	33.2 60	0.45 5
1	. .	I	. .	3	0.03	0.03 7	33.9 16	0.47 2
	* .		* .	3 5	0.07 8	0.10 2	37.0 98	0.37 2
	. .		. .	3	0.01 5	0.00 7	37.2 11	0.41 3

The new correlogram has by far less statistically significant terms, therefore we should search for a sample of ARIMA (3,1,3) type, and even if we take into account how separate are the significant terms, it is possible that this sample be actually ARIMA (1,1,1).

### 2.2 Sample estimation

The stage of sample estimation includes the effective use of data to do parameter inferences according to the soundness of the sample. In order to estimate parameters the method of maximum probability also known as the method of maximum likelihood or the method of the least squares can be used.

By using least squares, we have estimated the following model in Eviews:

d(p\_rrc) c ar(1) ar(2) ar(3) ma(1) ma(2) ma(3)

# Dependent Variable: D(P\_RRC)

Method: Least Squares

Sample (adjusted): 1/06/2006

11/15/2007

Included observations: 485 after

adjustments

Convergence achieved after 78 iterations Backcast: 1/03/2006 1/05/2006

Variable	Coeffic ient		t- Statistic	Prob.
С	05	3.50E- 05	9	0.1732
AR(1)	6	0.21729 4	7	0.0598
AR(2)	6 0.54719	0.25961 4 0.17166	1 3.18752	0.9079
AR(3)	5	8	7	0.0015
MA(1)	0.30983 5	0.20375 6	1.52061 9	0.1290
MA(2)	0.02281 2 -	0.22453 8	0.10159 7 -	0.9191
MA(3)	0.65674 5	0.15506 7	4.23522	0.0000
R-squared Adjusted R-squared	8	Mean do var S.D. do var	_	05
S.E. or regression	f 0.00267 0	Akaike criterion		- 8.99894 5
Sum squared resid Log likelihood	10.00340 8 2189.24 4	criterion	l	8.93855 5 3.39023 0
Durbin- Watson stat	1.97923 3	Prob(F-s	statistic)	0.00276 6
Inverted AR Roots Inverted MA Roots	.97	-	28+.70 3474i	
Tolzina into a		at the te	ma AD	(2) and

Taking into account that the terms AR (2) and MA (2) are statistically non-significant, we reestimate the sample without these:

Dependent Variable:

D(P RRC)

Method: Least Squares

Sample (adjusted): 1/06/2006

11/15/2007

Included observations: 485 after

adjustments

Convergence achieved after 56

iterations

Backcast: 1/03/2006 1/05/2006

Variable	Coeffic ient	Std. Error	t- Statisti c	Prob.
С	4.84E- 05 0.36400	05	87	0.2021
AR(1)	3	25	49	0.0054
AR(3)	0.56656	0.1250 66	4.5301 14	0.0000
MA(1)	0.29199 5	0.1138 41	2.5649 37	0.0106
MA(3)	0.69692 5	0.1145 52	6.0839 24	0.0000
R-squared Adjusted R squared	0.03755 2 -0.02953 2	depend	ent var ependent	
S.E. or regression	f 0.00266 9	Akaike criterio		9.00379 3
Sum squared resid	10.00342 0	Schwar criterio		8.96065 8
Log likelihood	2188.42 0	F-statis	tic	4.68211 6
Durbin- Watson stat	1.92713 4	Prob(F- statistic		0.00102 6
Inverted AF Roots Inverted MA Roots	97	- .30+.70 i 35- .76i	3070 35+.70	

In this sample, all coefficients except the constant are statistically significant.

### 2.3 Sample Checking

This last stage of the Box-Jenkins methodology is at least equally important as identification or estimate stage. The purpose is seeing in what extent the sample built complies with the available observations dealing with the stochastic process studied.

The stage implies testing the sample adjusted in its relation with data in order to discover the inadequacies of the sample and to obtain its improvement.

Taking into account that we have estimated an ARIMA(3,1,3) sample, we are in the first instance interested in knowing if we have eliminated autocorrelation of residuals. The correlogram of residuals (in the object equation -> view -> residual tests -> correlogram Q statistic) proves that there are no more autoregressive statistically significant terms. For verify this assumption we can used the Breusch-Godfrey test.

Breusch-Godfrey Serial Correlation LM Test:

	0.7965			0.45149
F-statistic	14	Prob. F(2,4	478)	6
Obs*R-	1.4132	Prob.	Chi-	0.49329
squared	81	Square(2)		9

**Test Equation:** 

Dependent Variable: RESID

Method: Least Squares

Sample: 1/06/2006 11/15/2007 Included observations: 485

Presample missing value lagged residuals set to zero.

Variable	Coeffic ient		t- Statistic	Prob.
С			0.00333 3	0.9973
AR(1)	9	701	0.45227 6 0.42149	0.6513
AR(3)	5	800	7 0.28572	0.6736
MA(1)	6	551	2	0.7752
MA(3)	1	752	0.28944 2 1.03552	0.7724
RESID(-1)	6	024	5	0.3009
RESID(-2)			0.26159	0 7937
TESID( 2)	=		=	=

R-squared	0.00291 4	Mean dependent var	5.37E- 05
Adjusted Resquared	- 0.00960 2	S.D. dependent var	0.00265 8
S.E. or regression	f 0.00267 0	Akaike info criterion	8.99887 3
Sum squared resid Log likelihood Durbin- Watson stat	9 2189.22 7	Schwarz criterion F-statistic Prob(F- statistic)	8.93848 3 0.23282 6 0.96581 3

The assumption can be accepted. Nevertheless, residuals are relatively far from normality, with both excess kurtosis and skewness positive (figure 2).

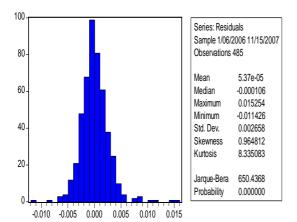


Figure 2 – The residual distribution

The test of double residual autocorrelation (squared residuals) also suggests that the heteroskedasticity hypothesis is not verified, and the ARIMA (3,1,3) sample should be estimated with a ARCH sample for variant, not at all simple least squares.

If we estimate the ARIMA (3,1,3) sample by means of a GARCH (1,1) sample for a variant, results are more encouraging:

### **Dependent Variable: D(P\_RRC)**

Method: ML - ARCH (Marquardt) - Normal distribution

Sample (adjusted): 1/06/2006 11/15/2007 Included observations: 485 after adjustments Convergence achieved after 72 iterations MA backcast: OFF (Roots of MA process too large), Variance backcast: ON

GARCH = C(6) + C(7)\*RESID(-1)^2 + C(8)\*GARCH(-1)

C(8)*GARCH(-1)				
	Coeffic ient	Std. Error	z- Statisti c	Prob.
С	-1.18E- 06	1.42E- 05	- 0.0831 07	0.933
AR(1)	3	4.52E- 05 0.00012	89	0
AR(3)	4	8 0.00047	57	0
MA(1)	9	3	198.93 06	0.000
MA(3)	0.88369 4	0.00014 9	5921.5 44	0.000
	Variance Equation			
С	07 0.25628	1.75E- 07 0.04633	19 5.5315	5 0.000
RESID(-1)^2 GARCH(-1)	3 0.68098 5	1 0.05683 4	70 11.982 00	0.000 0
R-squared Adjusted R-squared	0.07611 8 -0.06256 0	depende		-4.33E 05 0.0027 10
S.E. or regression	f 0.00262 3	Akaike criterion		9.2769 88
Sum squared resid Log	3 2257.67	criterion	l	9.2079 71 5.6142
likelihood Durbin- Watson stat	1.74531	F-statist Prob(F-s		69 0.0000 03
Inverted AR Roots Inverted MA Roots	.83	-	5879 6182 proce	i

Now, the residuals distribution is presented in figure 3

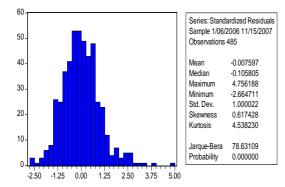


Figure 3 – The residual distribution

### 3. CONCLUSIONS

ARIMA(3,1,3) sample, possibly with a GARCH (1,1) sample for the variant of residuals, adequately describes the structure of autocorrelation in the field of Rompetrol share prices.

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